Distribution of Values of Real Quadratic Zeta Functions

Joshua Holden

Abstract. The author has previously extended the theory of regular and irregular primes to the setting of arbitrary totally real number fields. It has been conjectured that the Bernoulli numbers, or alternatively the values of the Riemann zeta function at odd negative integers, are evenly distributed modulo $p$ for every $p$. This is the basis of a well-known heuristic, given by Siegel in [16], for estimating the frequency of irregular primes. So far, analyses have shown that if $\mathbb{Q}(\sqrt{D})$ is a real quadratic field, then the values of the zeta function $\zeta_D(1 - 2m) = \zeta_{\mathbb{Q}(\sqrt{D})}(1 - 2m)$ at negative odd integers are also distributed as expected modulo $p$ for any $p$. However, it has proven to be very computationally intensive to calculate these numbers for large values of $m$. In this paper, we present the alternative of computing $\zeta_D(1 - 2m)$ for a fixed value of $D$ and a large number of different $m$.

1. Introduction, Conjectures, and Previous Results

Siegel, in [16], conjectured that the numerators of the Bernoulli numbers $B_{2m}$ were evenly distributed modulo $p$ for any odd prime $p$. This hypothesis was used by Johnson ([14]) and independently by Wooldridge ([20]) to predict the density of primes with a given index of irregularity, that is such that $p$ divides a given number of the Bernoulli numbers $B_2, \ldots, B_{p-3}$. It also comes in handy for predicting many other values that are related to irregular primes, such as the order of magnitude of the first prime of a given index of irregularity. (See, for example, [18].)

Since $B_{2m} = -\zeta(1 - 2m)(2m)$, it is equivalent to say that the values of $\zeta(1 - 2m)$ are evenly distributed modulo $p$, where $\zeta(s)$ is the Riemann zeta function. Of course, the Riemann zeta function can be generalized to any number field $k$ to get a zeta function $\zeta_k(s)$ associated with the number field $k$. If $k$ is a totally real number field, there are a number of situations where the values $\zeta_k(s)$ can be seen as analogous to the Bernoulli numbers, including generalizations of Kummer’s Criterion for predicting when $p$ divides the class number of the $p$-th cyclotomic field and cases where the equation $x^p + y^p = z^p$ of Fermat’s Last Theorem can be shown not to have solutions in $k$. (See for example, [5, 6, 7], [10], [9], [10], [11], and [15].)
Little or no progress has been made on proving Siegel’s hypothesis, but a great deal of data has been collected, especially in regard to the prediction of Johnson and Wooldridge. Specifically, this prediction says that as \( p \to \infty \), the probability that \( p \) has index of irregularity \( r \) goes to

\[
\left( \frac{1}{2} \right)^r e^{-1/2} \frac{1}{r!}.
\]

(In addition to the original sources, the details may be found in Section 5.3 of [19].) Note that this prediction does not rely on the full strength of Siegel’s hypothesis, but merely on the weaker hypothesis that the Bernoulli numbers are \( 0 \mod p \) with probability \( 1/p \). The investigations focused on in this paper relate only to predictions about indices of irregularity based on this weaker hypothesis.

Wagstaff, in [18], computed \( u_r(x) \), the fraction of primes not exceeding \( x \) with index \( r \) of irregularity for each \( r \) between 0 and 2 and for all \( r \geq 3 \) grouped together, and compared this distribution to the predicted distribution for each multiple \( x \) of 1000 up to 125000. The result of the chi-squared test “fluctuated usually between 0.1 and 1.0 and had the value 0.29 at \( x = 125000 \). It was 0.03 at \( x = 8000 \)” [18]. These results correspond to significance levels of .992, .801, .962, and .999, respectively. (The significance levels used in this paper correspond roughly to the probability that the agreement between the observed results and the predicted results is not due to chance. Statisticians consider the threshold for considering a result to be not due to chance to be a significance level of .9 to .95. Since we are not actually conducting a valid statistical study in this paper, all of the statistical results should be taken with a very large grain of salt.)

Buhler, Crandall, Ernvall, and Metsänkylä hold the record for computations with irregular primes, having found all the irregular primes below four million as described in [2]. They do not seem to have done a chi-squared analysis, but they tabulate the values of \( u_r(x) \) for \( x = 4000000 \) and \( r \) between 0 and 7. A chi-squared test using the same methodology as before has the result 1.02, for a significance level of .796. Earlier, in [3], Buhler, Crandall, and Sompolski tabulated the same data for \( x = 1000000 \). The result of the same chi-squared test is 0.78, for a significance level of .854.

Unfortunately, the only way to collect data to test Siegel’s hypothesis is to investigate \( B_{2m} \) for larger and larger \( m \), which is very computationally intensive. (See [1] or [8] for details.)

However, in the more general number field case, there are many more dimensions to the problem. We start by restricting our attention to the case of \( k \) an abelian totally real number field. Then we know that

\[
\zeta_k(s) = \prod_{\chi \in \hat{G}} L(s, \chi)
\]

where \( \hat{G} \) is the character group of \( G = \text{Gal}(k/Q) \) and \( L(s, \chi) \) is the \( L \)-function associated with the character \( \chi \). Note that \( L(s, 1) = \zeta(s) \), so the Riemann zeta function is a factor of the zeta function for \( k \). (See [4], e.g., for more details.) Certainly it seems likely that for a fixed (totally real) number field \( k \) and character \( \chi \) the values of the numerator of \( L(1-2m, \chi) \) are evenly distributed modulo \( p \) as \( m \) varies. (It is known that these values are rational numbers.) We also hypothesize that these values for different \( \chi \) are independent, which implies that the numerators of \( \zeta_k(1-2m) \) are distributed modulo \( p \) like the product of \( |G| \) independent integer
variables, each of which is evenly distributed modulo $p$. We will refer to this as the “product distribution”, for lack of a better term. However, it also is reasonable to conjecture that for a fixed $m$ the values of $\zeta_k(1 - 2m)$ are distributed according to the product distribution modulo $p$ as $k$ varies. More precisely, if we fix $m$ and the degree of $k$ we expect the values to be distributed according to the product distribution modulo $p$ as the discriminant of $k$ varies. Alternatively, if we fix $m$ and the discriminant of $k$ we expect the values to be distributed according to the product distribution modulo $p$ as the degree varies.

In this paper we will be considering the former situation, with the degree fixed at 2, making $k$ a real quadratic field. We let $k = \mathbb{Q}(\sqrt{D})$ (where $D$ is the discriminant of $k$) and $\zeta_D(s) = \zeta_k(s)$. In this case

$$\zeta_D(s) = L(s,1)\zeta(s) = L(s,\chi)$$

where $\chi(s) = \left( \frac{D}{\cdot} \right)$, the Kronecker symbol, where appropriate.

We make the following definitions:

**Definition 1.** Let $k = \mathbb{Q}(\sqrt{D})$ be a real quadratic number field with discriminant $D$. We say that an odd prime $p$ is $k$-regular (or $D$-regular) if $p$ is relatively prime to $\zeta_k(1 - 2m)$ for all integers $m$ such that $2 \leq 2m \leq \delta - 2$ and also $p$ is relatively prime to $p\zeta_k(1 - \delta)$, where $\delta = p - 1$ unless $D = p$, in which case $\delta = (p - 1)/2$. The number of such zeta-values that are divisible by $p$ will be the index of $k$-irregularity (or index of $D$-irregularity) of $p$.

Further, we will say that $p$ is $\chi$-regular if $p$ is relatively prime to $L(1 - 2m, \chi)$ for all integers $m$ such that $2 \leq 2m \leq \delta - 2$ and also $p$ is relatively prime to $pL(1 - 2m, \chi)$. The number of such $L$-values that are divisible by $p$ will be the index of $\chi$-irregularity of $p$.

(See [11], [12] or [13] for an explanation of why the definition has exactly this form.)

Saying that the values of $\zeta_k(1 - 2m)$ are distributed according to the product distribution and that the values of $\zeta(1 - 2m)$ are evenly distributed is the same as saying that the values of $L(1 - 2m, \chi)$ are evenly distributed modulo $p$. Then we can make the same prediction about the indices of chi-irregularity that Johnson and Wooldridge made about the indices of irregularity in the rational case. We briefly investigated this issue in [11], where there are tables of the analogue of $u_r(x)$ (using the index of $\chi$-irregularity) for $x = 1000$, $r$ from 0 to 4, and $D = 5, 8, 12,$ and 13. The chi-squared test results are not included, but using the methodology discussed earlier they are 3.32, 1.74, 1.15, and 2.54. The corresponding significance levels are .345, .628, .765, and .469, respectively. We could total the values of (the analogue of) $u_r(x)$ for the four values of $D$ and compare them to the predicted values; we might expect that this would give us a better significance level because of the larger “sample size”. However, in this case the chi-squared result is 3.53 and the significance level is .316, which is worse than any of the results for the values of $D$ taken separately! We could instead average the values for the four values of $D$ and compare them to the predicted values; in this case the chi-squared result is 0.884 and the significance level is .829, which is quite good. However, it is not clear to us what the actual meaning of these averages are. Let the reader beware.

The papers [11] and [13] investigate some algorithms, based on formulas given by Siegel in [17], for calculating values of $\zeta_D(1 - 2m)$. We show there that while the best-known algorithms for calculating $\zeta_D(1 - 2m)$ letting $m$ vary take amortized
time polynomial in \( m \), it is possible to calculate \( \zeta_D(1-2m) \) with \( m \) fixed and \( D \) varying in amortized time subpolynomial in \( D \). We will use such an algorithm in the following.

2. New Results

In the course of testing the algorithms in [13], we collected more data in addition to that above. Table 1 shows the number of primes less than 5000 which have \( \chi \)-index of irregularity \( r \) for various values of \( r \) and \( D = 5 \). We compared the observed and predicted distributions, using the methodology above, for primes below \( x \) where \( x \) was 1000, 2000, 3000, 4000, and 5000, and found chi-squared values of 3.32, 5.03, 2.51, 1.73, and 2.10 and significance levels of .344, .170, .473, .630, and .552, respectively.

<table>
<thead>
<tr>
<th>( r )</th>
<th>number</th>
<th>predicted number</th>
<th>predicted fraction</th>
</tr>
</thead>
<tbody>
<tr>
<td>0</td>
<td>422</td>
<td>405.16</td>
<td>.606531</td>
</tr>
<tr>
<td>1</td>
<td>186</td>
<td>202.58</td>
<td>.303265</td>
</tr>
<tr>
<td>2</td>
<td>51</td>
<td>50.65</td>
<td>.075816</td>
</tr>
<tr>
<td>3</td>
<td>7</td>
<td>8.44</td>
<td>.012636</td>
</tr>
<tr>
<td>4</td>
<td>2</td>
<td>1.06</td>
<td>.001580</td>
</tr>
</tbody>
</table>

Table 1. Results for \( D = 5 \) and \( p < 5000 \)

Other data was obtained using the philosophy, described above, of computing the values of \( L(1-2m, \chi) \) for large numbers of \( D \) and relatively small values of \( m \). As in the discussion of \( D = 5, 8, 12, \text{and 13} \) above, we present both the total and the average across the different discriminants. Table 2 presents the data for all \( D < 5000 \) and \( p < 100 \). The chi-squared value for the totals is 81.1 and the significance level is .000. The chi-squared value for the averages is 0.053 and the significance level is .997. As before, the actual meaning of these numbers is not clear.

<table>
<thead>
<tr>
<th>( r )</th>
<th>total number</th>
<th>predicted total number</th>
<th>average number</th>
<th>predicted average number</th>
<th>predicted fraction</th>
</tr>
</thead>
<tbody>
<tr>
<td>0</td>
<td>21864</td>
<td>22068.01</td>
<td>14.42</td>
<td>14.56</td>
<td>.606531</td>
</tr>
<tr>
<td>1</td>
<td>11596</td>
<td>11034.01</td>
<td>7.65</td>
<td>7.28</td>
<td>.303265</td>
</tr>
<tr>
<td>2</td>
<td>2529</td>
<td>2758.50</td>
<td>1.67</td>
<td>1.82</td>
<td>.075816</td>
</tr>
<tr>
<td>3</td>
<td>347</td>
<td>459.75</td>
<td>0.23</td>
<td>0.30</td>
<td>.012636</td>
</tr>
<tr>
<td>4</td>
<td>41</td>
<td>57.47</td>
<td>0.03</td>
<td>0.04</td>
<td>.001580</td>
</tr>
<tr>
<td>5</td>
<td>7</td>
<td>5.75</td>
<td>0.005</td>
<td>0.004</td>
<td>.000158</td>
</tr>
</tbody>
</table>

Table 2. Results for \( D < 5000 \) and \( p < 100 \)

Finally, we computed the values of \( L(-1, \chi) \) and \( L(-3, \chi) \) for the 303957 discriminants \( D \) less than one million, and calculated the indices of \( \chi \)-irregularity for the primes 3 and 5. In this case, since the \( p \) involved are very small, we computed the expected number of primes with given index of irregularity directly from Siegel’s hypothesis, rather than using the limit as \( p \to \infty \). This takes into account, for instance, the fact that for \( p = 3 \) the index of irregularity cannot be
more than 1 and for \( p = 5 \) it cannot be more than 2. The results are shown in Table 3. The chi-squared value for the totals is 24636 and the significance level is .000. (In this case the categories used for the chi-squared test are \( r = 0 \), \( r = 1 \), and \( r = 2 \).) The chi-squared value for the averages is 0.081 and the significance level is .960. Nevertheless, in view of the large amount of data there seems to be a significant discrepancy between the predicted results and the observed results, even when the averages are considered. Coupled with our doubts about the meaning of the averaged numbers, this leads us to believe that something other than a straightforward extension of Siegel’s hypothesis is influencing the distribution of the values of \( L(1 - 2m, \chi) \) as the discriminant varies. We hope to make the nature of this influence clearer in the future.

Table 3. Results for \( D < 1000000 \) and \( p = 3 \) or \( p = 5 \)

<table>
<thead>
<tr>
<th>( r )</th>
<th>total number</th>
<th>predicted total number</th>
<th>average number</th>
<th>predicted average number</th>
<th>predicted fraction</th>
</tr>
</thead>
<tbody>
<tr>
<td>0</td>
<td>338966</td>
<td>397170.48</td>
<td>1.115177</td>
<td>1.306667</td>
<td>0.653333</td>
</tr>
<tr>
<td>1</td>
<td>252832</td>
<td>198585.24</td>
<td>0.831802</td>
<td>0.653333</td>
<td>0.326667</td>
</tr>
<tr>
<td>2</td>
<td>16116</td>
<td>12158.28</td>
<td>0.053021</td>
<td>0.040000</td>
<td>0.020000</td>
</tr>
</tbody>
</table>

3. Other Conjectures and Future Work

A number of other conjectures about irregular primes could be extended to the setting which we have presented. One such, mentioned by Wagstaff in [18] without attribution, is that the irregular primes are evenly distributed across the possible residue classes modulo \( n \) for every positive integer \( n \). To be exact, one expects that the ratio of irregular primes in a residue class to all odd primes in the class to be the same for each possible residue class. Wagstaff investigated this for primes below 125000 and \( 3 \leq n \leq 37 \) (and also \( n = 59 \) and 67) and gave tables for \( n = 3, 4, \) and 5. He does not give the chi-squared results, but they are 0.107, 0.060, and 2.420, respectively, with significance levels of .744, .806, and .490. We conjecture that the \( \chi \)-irregular primes are also distributed evenly, and intend to investigate this using the data we have collected.

Wagstaff also mentions a conjecture due to Wooldridge about the distribution of the numbers \( 2k/p \) for which \( (p, 2k) \) is an irregular pair; i.e. \( p \) divides \( B_{2k} \). Wooldridge conjectures that these numbers have a uniform distribution in the interval \((0, 1)\). We also intend to investigate whether these numbers are uniformly distributed for \( \chi \)-irregular pairs.

Two other conjectures about irregular primes which are widely believed cannot really be checked using statistical methods. It is thought that there are primes with arbitrarily large indices of irregularity; the largest observed so far is 7, for \( p = 3238481 \) (see [2]). The largest index of \( \chi \)-irregularity yet observed is 5, which occurred 7 times among the primes less than 100 using discriminants less than 5000. It is also likely that there are Bernoulli numbers divisible by arbitrarily large powers of \( p \). Although as yet \( p^2 \) has not been seen to divide \( B_{2m} \), as Washington says, “there does not seem to be any reason to believe this in general” [19]. On the other hand, powers as large as \( 3^7 \) have been observed to divide \( L(1 - 2m, \chi) \); this happens at \( D = 3869 \) and again at \( D = 3937 \).
Studies such as [2] also frequently investigate applications such as Fermat’s Last Theorem, Vandiver’s conjecture, and Iwasawa’s cyclotomic invariants. All of these could be modified to the situation explored here. Additional data would have to be collected in order to study these applications; we are not yet certain whether new algorithms would have to be implemented.

References


Department of Mathematics, Duke University, Durham, NC 27708, USA
E-mail address: holden@math.duke.edu